

# Probability Theory

## Solutions #10

10.4 If  $X$  is log-normal with parameters  $(m, \sigma)$  then it has the same distribution as  $e^{\sigma Z + m}$  where  $Z$  is standard normal. Then  $CX^\alpha$  has the same distribution as  $e^{\alpha\sigma + \alpha m + \log C}$  thus it is log-normal with parameters  $(\alpha\sigma, \alpha m + \log C)$ .

10.7 You can find the density functions by several methods. You can use the formula we had in class:  $g(y) = \sum_{x \in \Psi^{-1}(y)} \frac{x}{|\Psi'(x)|}$ , but you have to remember that  $\Psi^{-1}(y)$  is a set and if  $\Psi$  is not one-to-one

then it will have more than one elements. (E.g. if  $\Psi(x) = x^2$ .) You can also find the density by first determining the distribution function and then differentiating it.

(a) We will calculate the distribution function of  $X$  for  $x \in [0, 1]$ :

$$F(x) = \mathbf{P}(\xi^2 < x) = \mathbf{P}(-\sqrt{x} < \xi < \sqrt{x}) = \sqrt{x}.$$

Thus  $f(x) = \frac{1}{2\sqrt{x}}$  for  $x \in [0, 1]$  and 0 otherwise.

For the density of  $Y$  it is easier to use the formula, since  $\tan(\pi/2x)$  is invertible if  $x \in (-1, 1)$ . The density will be  $g(y) = \frac{1}{2} \frac{2}{\pi} (\arctan y)' = \frac{1}{\pi} \frac{1}{1+y^2}$ .

(b) Set  $h(x) = \lambda e^{-\lambda x}$  for  $x > 0$  and 0 otherwise (this is the density of  $\xi$ ). To get the density of  $X$  we only need a linear transformation: the density will be  $\frac{1}{3} h(\frac{x-2}{3})$ .

Since the function  $\sqrt{y}$  is invertible for  $y > 0$  we get  $g(y) = h(y^2)2y$  for the density of  $Y$ .

(c) If  $x > 0$  then the distribution function of  $X$  is

$$\mathbf{P}(\xi^{-2} < x) = \mathbf{P}(\xi^2 > 1/x) = \mathbf{P}(\xi > \sqrt{1/x} \text{ or } \mathbf{P}(\xi < -\sqrt{1/x})) = 2 - 2\Phi(\sqrt{1/x}).$$

By differentiation  $f(x) = \phi(\sqrt{1/x})\sqrt{1/x^3} = \frac{1}{2\pi x^3} e^{-1/2x}$  if  $x > 0$  and 0 otherwise.

10.12 It is enough to prove the inequality for  $\sigma = 1$ , the general case follows from scaling. By differentiating all three terms, and also checking that there limits are 0 when  $x \rightarrow \infty$  we get that

$$\begin{aligned} \frac{1}{\sqrt{2\pi}} e^{-(x^2/2)} \left( \frac{1}{x} - \frac{1}{x^3} \right) &= \int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-(y^2/2)} \left( 1 - \frac{3}{y^4} \right) dy \\ 1 - \Phi(x) &= \int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-(y^2/2)} dy \\ \frac{1}{\sqrt{2\pi}} e^{-(x^2/2)} \frac{1}{x} &= \int_x^\infty \frac{1}{\sqrt{2\pi}} e^{-(y^2/2)} \left( 1 + \frac{1}{y^2} \right) dy \end{aligned}$$

Comparing the functions under the integrals we get the needed inequalities.

10.13 We need to find  $\mathbf{P}(55 < X < 65)$  in the cases where  $X$  is normal/log-normal with  $\mathbf{E}X = 60$  and  $\mathbf{Var}X = 3^2$ .

(a) If  $X$  is normal then it has to be  $N(60, 3)$  which means that  $\frac{X-60}{3}$  is standard normal. Thus

$$\mathbf{P}(55 < X < 65) = \mathbf{P}(-5/3 < \frac{X-60}{3} < 5/3) = 2\Phi(5/3) - 1.$$

(b) This needs a bit more calculation. If  $X$  is log-normal with parameters  $(m, \sigma)$ , then  $\mathbf{E}X = e^{\sigma^2/2+m}$  and  $\mathbf{Var}X = e^{2\sigma^2+2m} - e^{\sigma^2+2m}$ . From these the values of  $m$  and  $\sigma$  may be easily computed. (They will not be equal to 60 and 3!) If  $X$  is log-normal with parameters  $(m, \sigma)$  then  $\frac{\log X - m}{\sigma}$  is standard normal. Then, by similar calculations as in the previous part,  $\mathbf{P}(55 < X < 65) = \Phi\left(\frac{\log 65 - m}{\sigma}\right) - \Phi\left(\frac{\log 55 - m}{\sigma}\right)$ .

10.14 (a) Observe that in all three cases we have to find the expectation of a  $f(X)$  where  $f(\cdot)$  is anti-symmetric:  $f(-x) = -f(x)$ . Thus by the symmetry of the density function of the standard normal distribution all three expectations will be 0.

(b) You need to use the identities  $\cos x = \frac{e^{ix} + e^{-ix}}{2}$ ,  $\sin x = \frac{e^{ix} - e^{-ix}}{2}$  (where  $i = \sqrt{-1}$ ), and also that the formula for the momentum-generating function of the standard normal also works for complex values:  $\mathbf{E}e^{itX} = e^{-t^2/2}$  (you need some complex analysis to verify this properly). From these you can compute all these expectations and variances, e.g.  $\mathbf{E} \cos X = \mathbf{E} \frac{e^{iX} + e^{-iX}}{2} = e^{-1/2}$  (the others may be computed similarly).

Since you needed complex analysis for this part, I gave full credit if you did part (a) of the problem.