

On the speed of convergence of Markov chain Monte Carlo methods

Outline

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- Upper and lower bounds for relaxation times

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- Myth busting
- More negative results
- Some hope
- Or proving that there's no hope...

Markov chain Monte Carlo

“Almost all Markov chains can be modified to converge to a prescribed distribution”

Primer Markov chain with transition probabilities $T(Y|X)$, aperiodic, irreducible and for all X and Y ,

$$T(Y|X) \neq 0 \implies T(X|Y) \neq 0$$

Metropolis–Hastings algorithm:

Step 1. (proposal) Propose a $Y \sim T(\cdot|X_n)$

Step 2. (acceptance) Accept Y with probability

$$\min \left\{ 1, \frac{\pi(Y)T(X_n|Y)}{\pi(X)T(Y|X_n)} \right\}$$

Mixing of Markov chains

“Mixing correlates with the second largest eigenvalue”

SLEM: Second largest eigenvalue modulus

$$\rho = \min\{\lambda_2, |\lambda_r|\}$$

Relaxation time

$$\tau_{rel} := \frac{1}{1 - \lambda_2}$$

$$\tau_i(\varepsilon) := \min \{k_0 : \forall k > k_0, d_v(P^k \delta_i, \pi) \leq \varepsilon\}$$

Theorems (Aldous, 1982; Diaconis & Stroock, 1991)

$$\tau_i(\varepsilon) \leq \frac{1}{1 - \rho} \left(\ln \left(\frac{1}{\pi(i)} \right) + \ln \left(\frac{1}{\varepsilon} \right) \right)$$

$$\max_{i \in I} \tau_i(\varepsilon) \geq \frac{\rho}{2(1 - \rho)} \ln(1/\varepsilon)$$

Cheeger inequalities

"A Markov chain mixes slowly if and only if there is a bottleneck in it"

Ergodic flow:

$$F(S) := \sum_{i \in S, j \in i \setminus S} \pi(i) p_{ij}$$

Conductance:

$$\Phi := \min_{S \in I} \left\{ \frac{\sum_{i \in S, j \in i \setminus S} \pi(i) p_{ij}}{\pi(S)} \mid 0 < \pi(S) \leq \frac{1}{2} \right\}$$

Cheeger inequality:

$$1 - 2\Phi \leq \lambda_2 \leq 1 - \frac{\Phi^2}{2}$$

MCMC convergence

MCMC convergence

D_1 D_2 D_3 D_4 D_5 D_6 D_7 ... D_n

MCMC convergence



MCMC convergence



Fast convergence

$\exists k \forall D_n (|D_n| \leq n)$ the series of $1/(n^k(1 - \lambda_{n,2}))$ built from the $\lambda_{n,2}$'s of the Markov chains M_n defined by the algorithms converges to 0.

MCMC convergence



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Slow convergence

$\exists D_n (|D_n| \leq n)$ s.t. $\forall k$ the series of $1/(n^k(1 - \lambda_{n,2}))$ built from the $\lambda_{n,2}$'s of the Markov chains M_n defined by the algorithms does not converge to 0.

Fast converging Markov chains

FPRAS:

(Fully Polynomial Randomized Approximation Scheme)

given $x \in \Sigma^*$, $\varepsilon > 0$, $0 < \delta < 1$:

$$\Pr((1 - \varepsilon)f(x) \leq OUT \leq (1 + \varepsilon)f(x)) \geq 1 - \delta$$

Algorithm polynomial in $|x|$, $1/\varepsilon$ and $-\log(\delta)$

#P: Hard counting problem

Fast converging Markov chains

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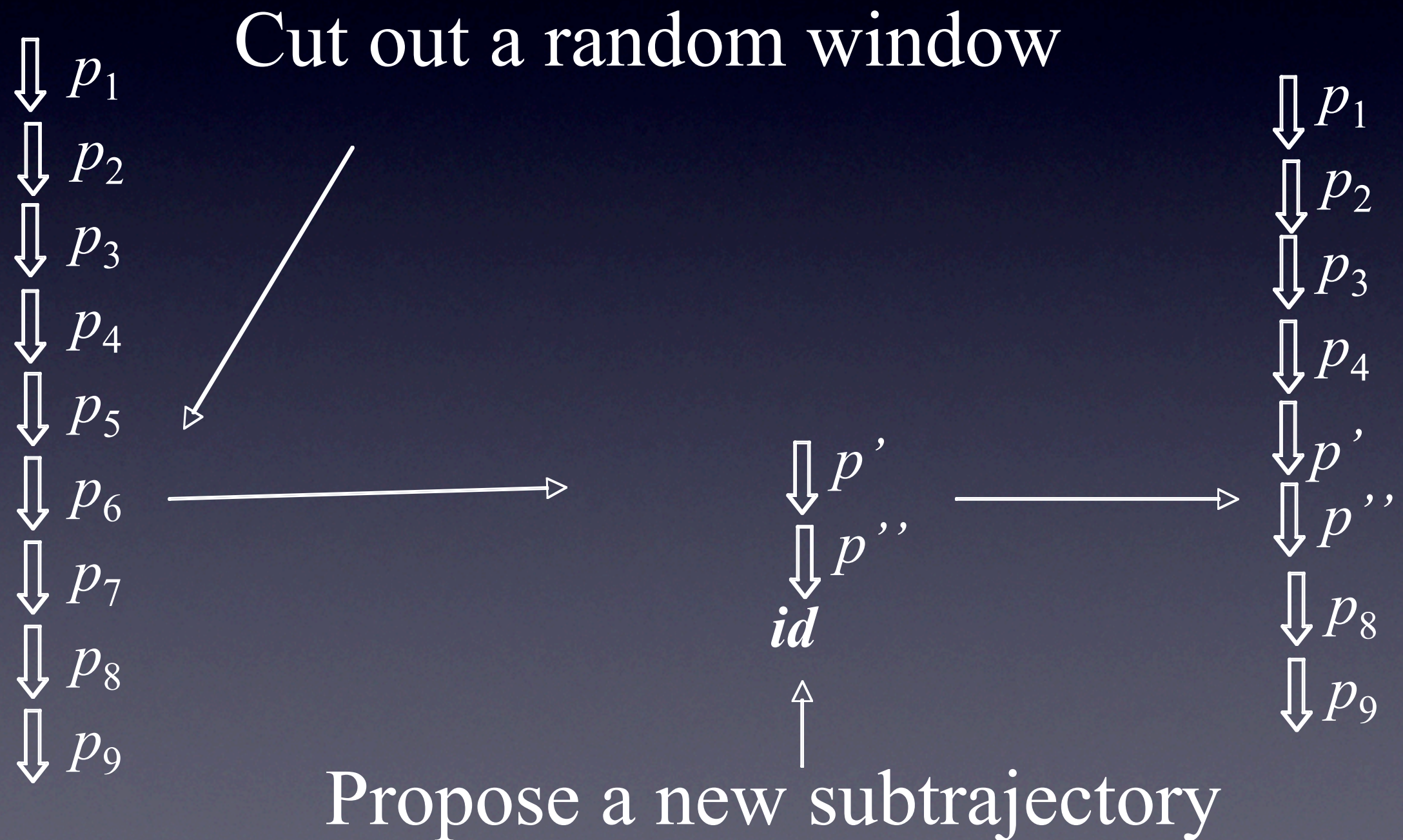
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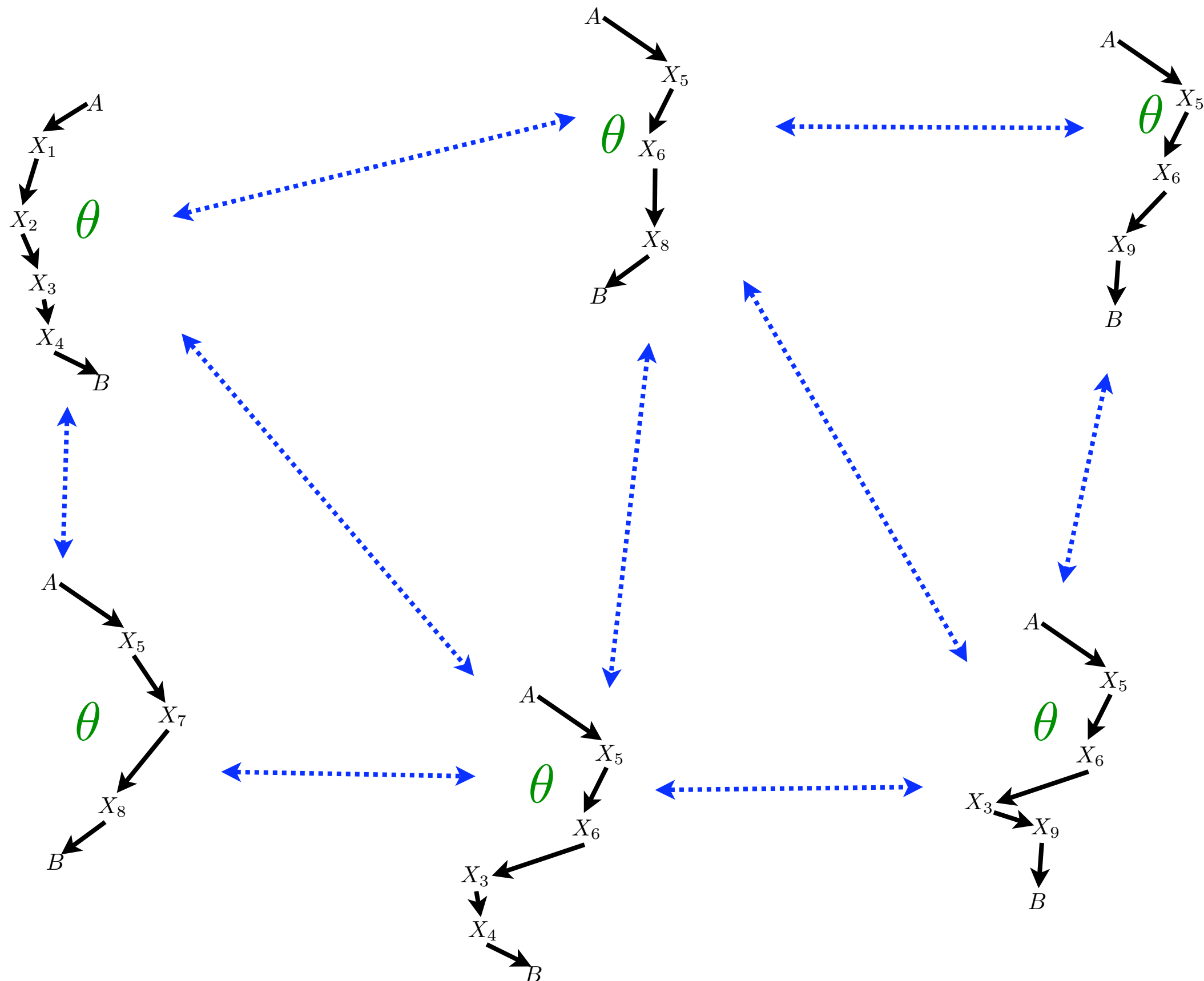
#P: Hard counting problem

It is known that the intersection of #P and FPRAS is not empty. For example linear extension of partially ordered sets.

ParIS: Partial Importance Sampling



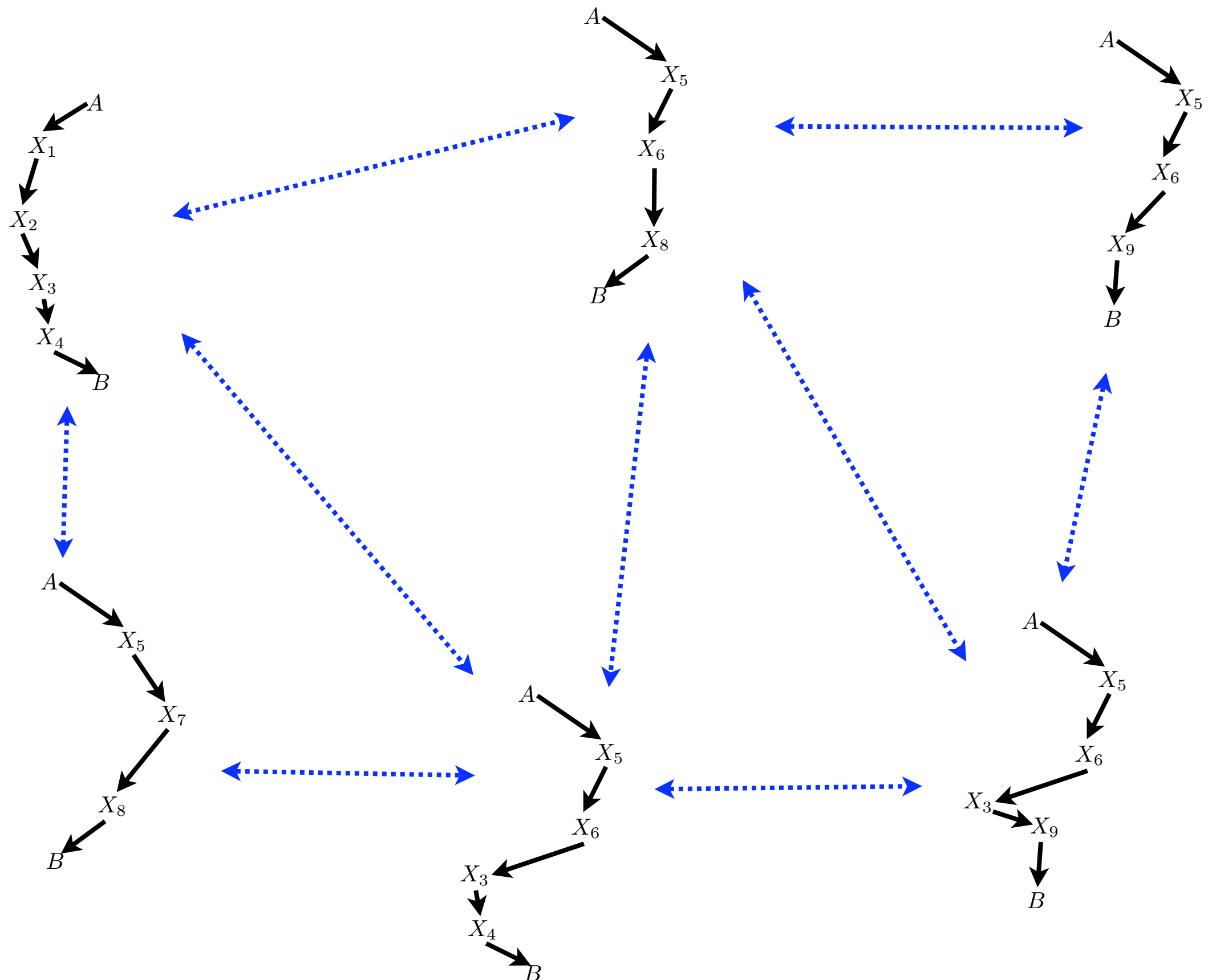
Random walk on random walks



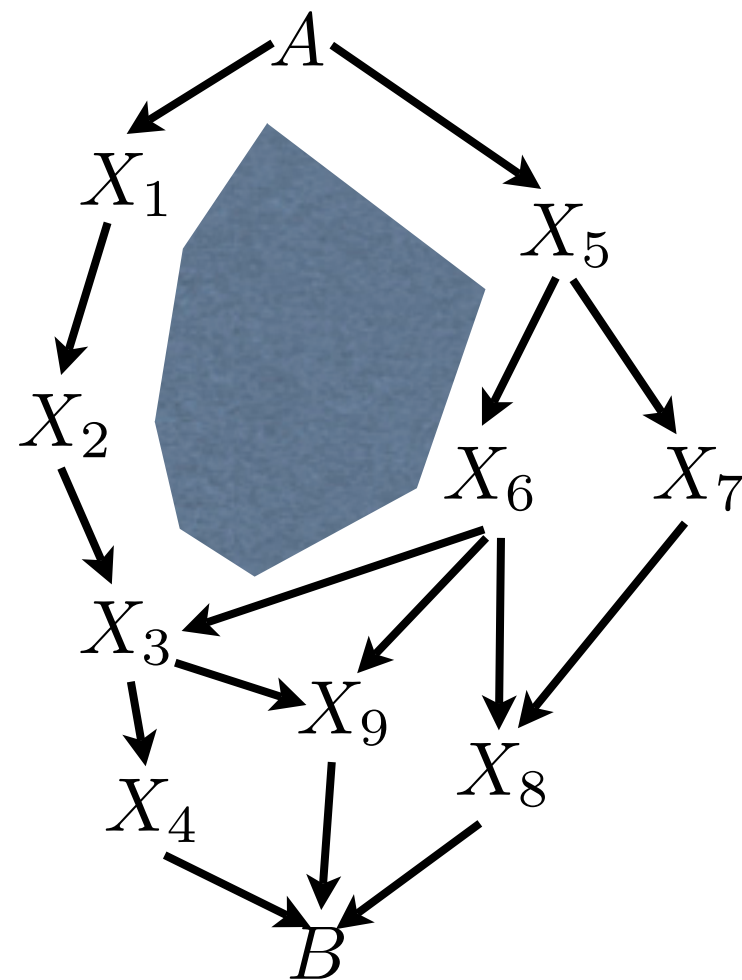
Examples for RWoRWs

- Genome rearrangement paths
- Ancestral states in the infinite site recombination model
- Sequence alignment
- etc.

Big islands

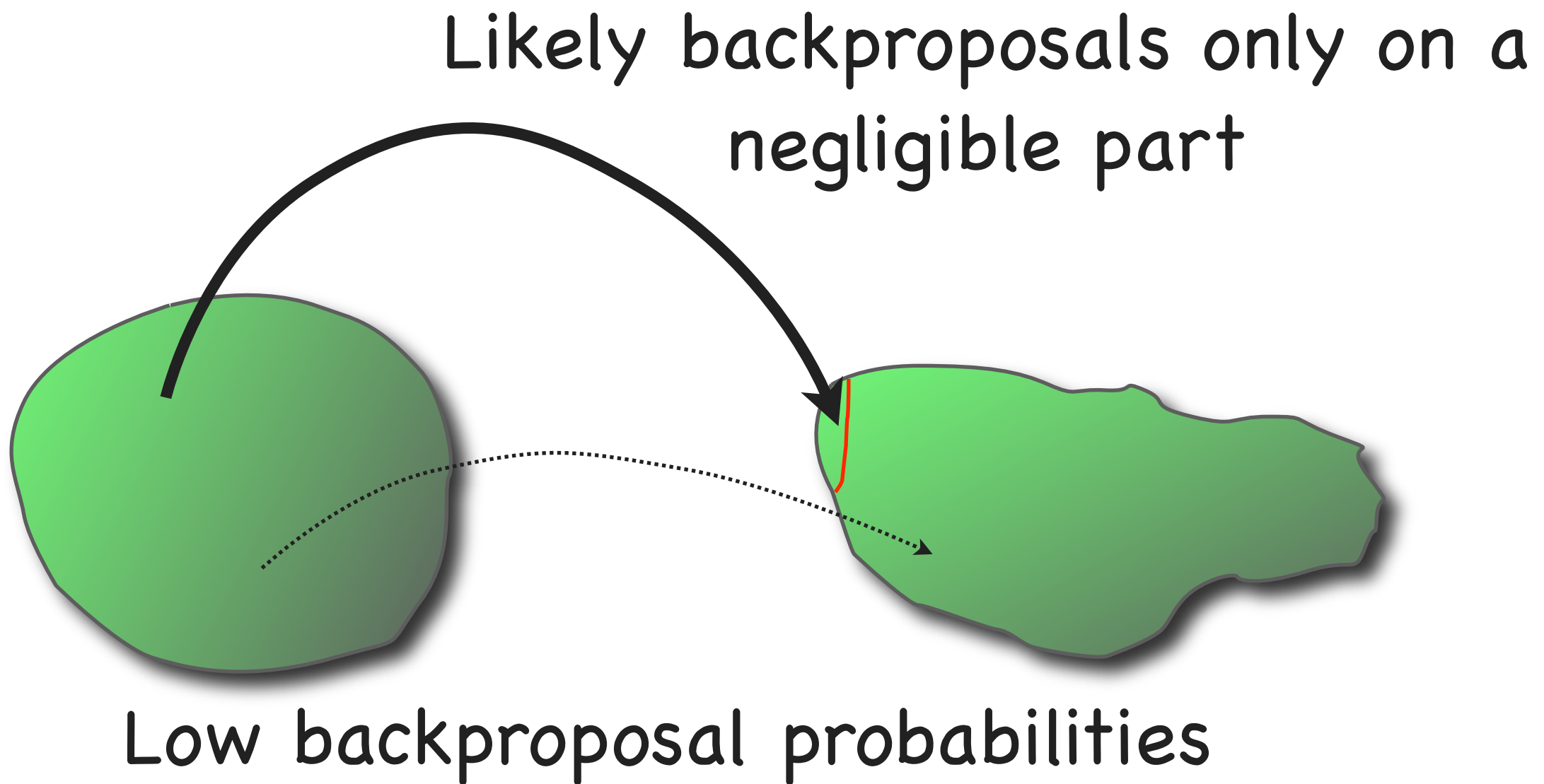


Big islands



ParIS mixes slowly on minimal reversal sorting paths

Miklós, Mélykúti, Swenson, manuscript submitted

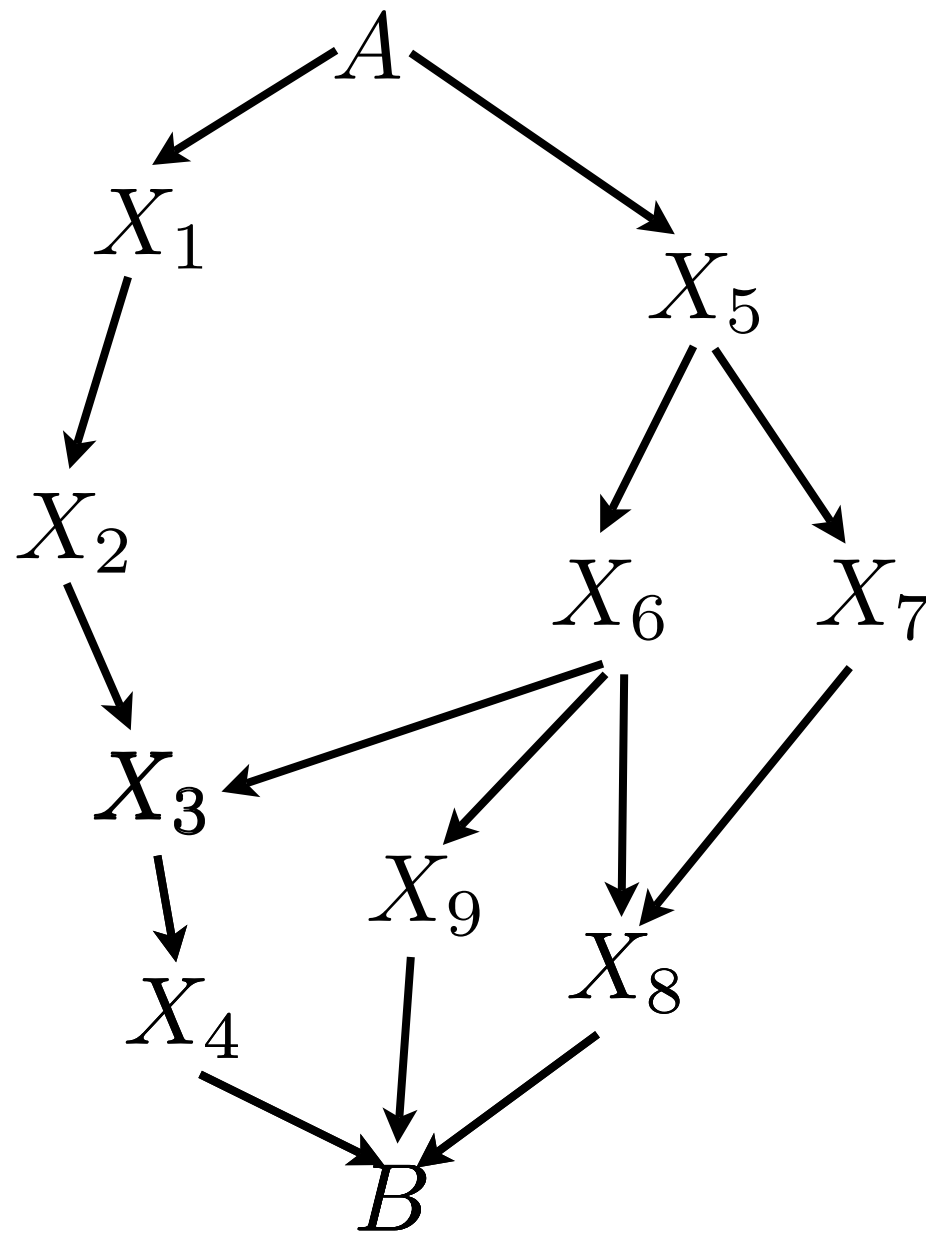




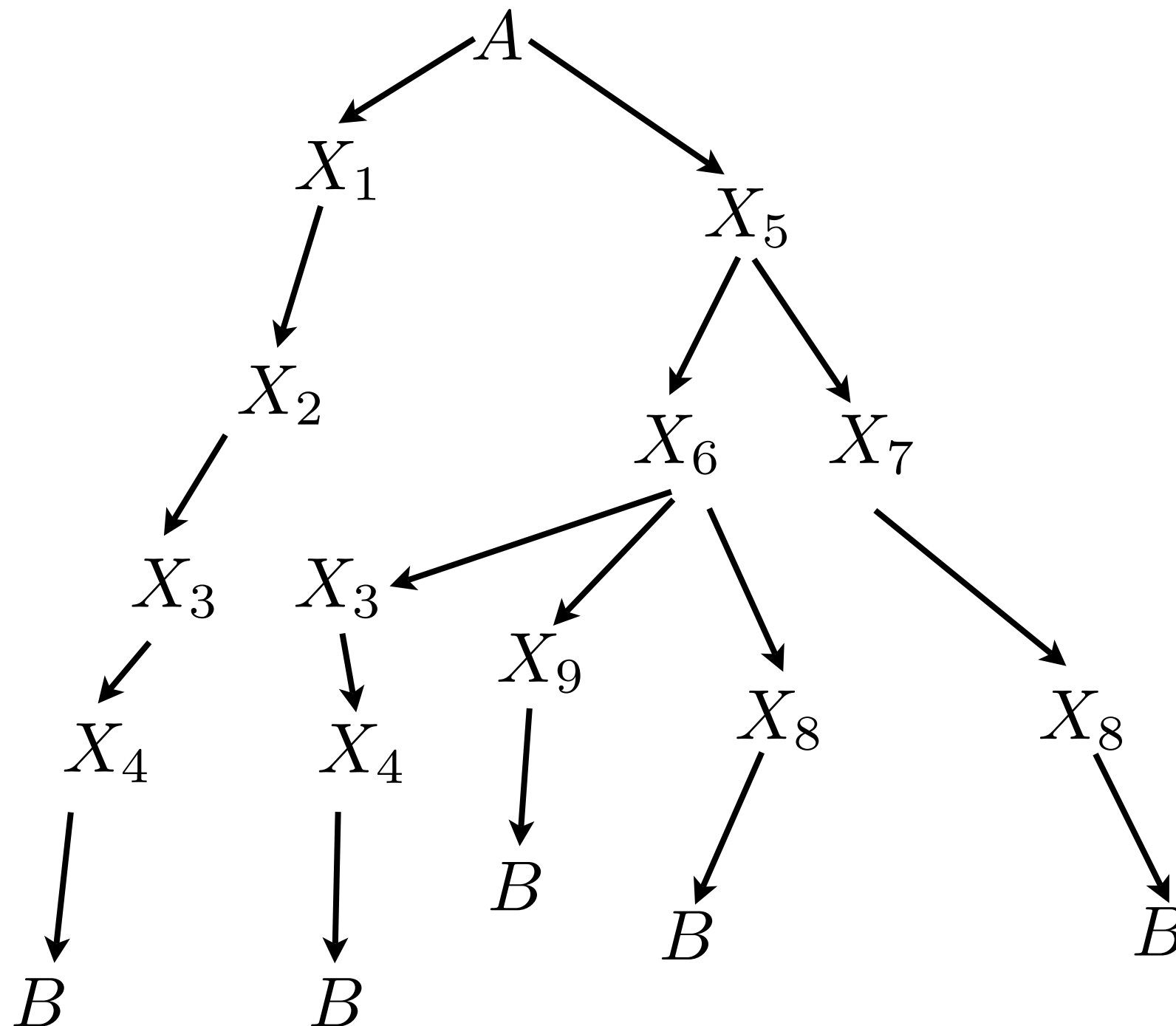
Myth I.

It is worth jumping big

Random walks as decision trees



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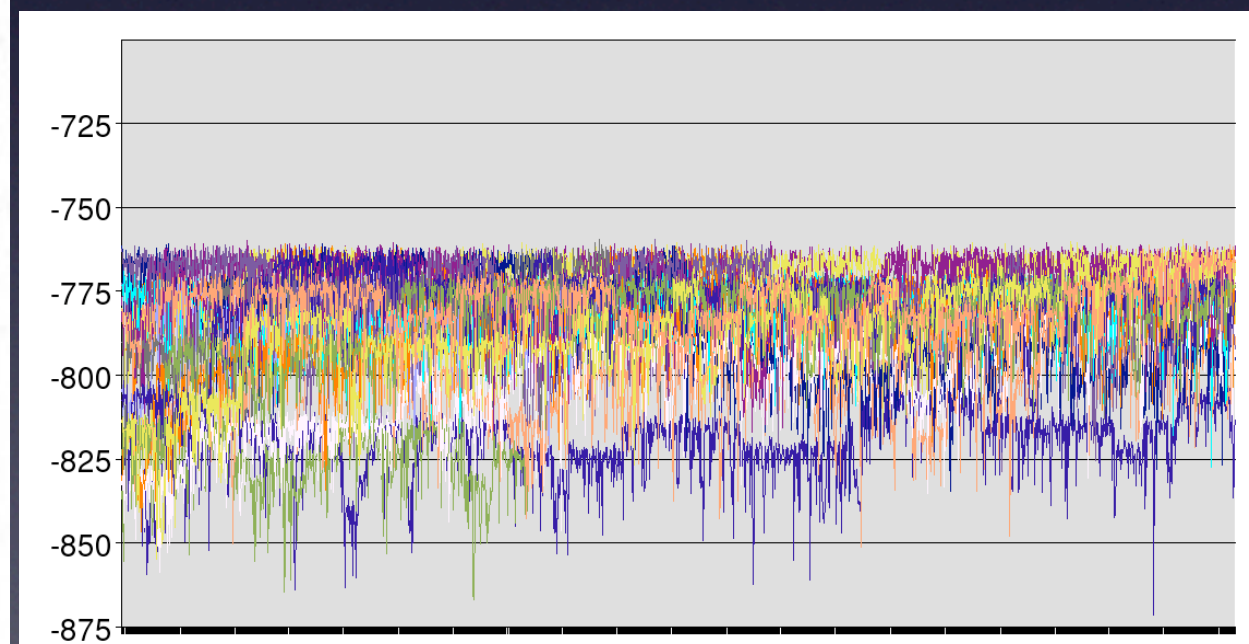
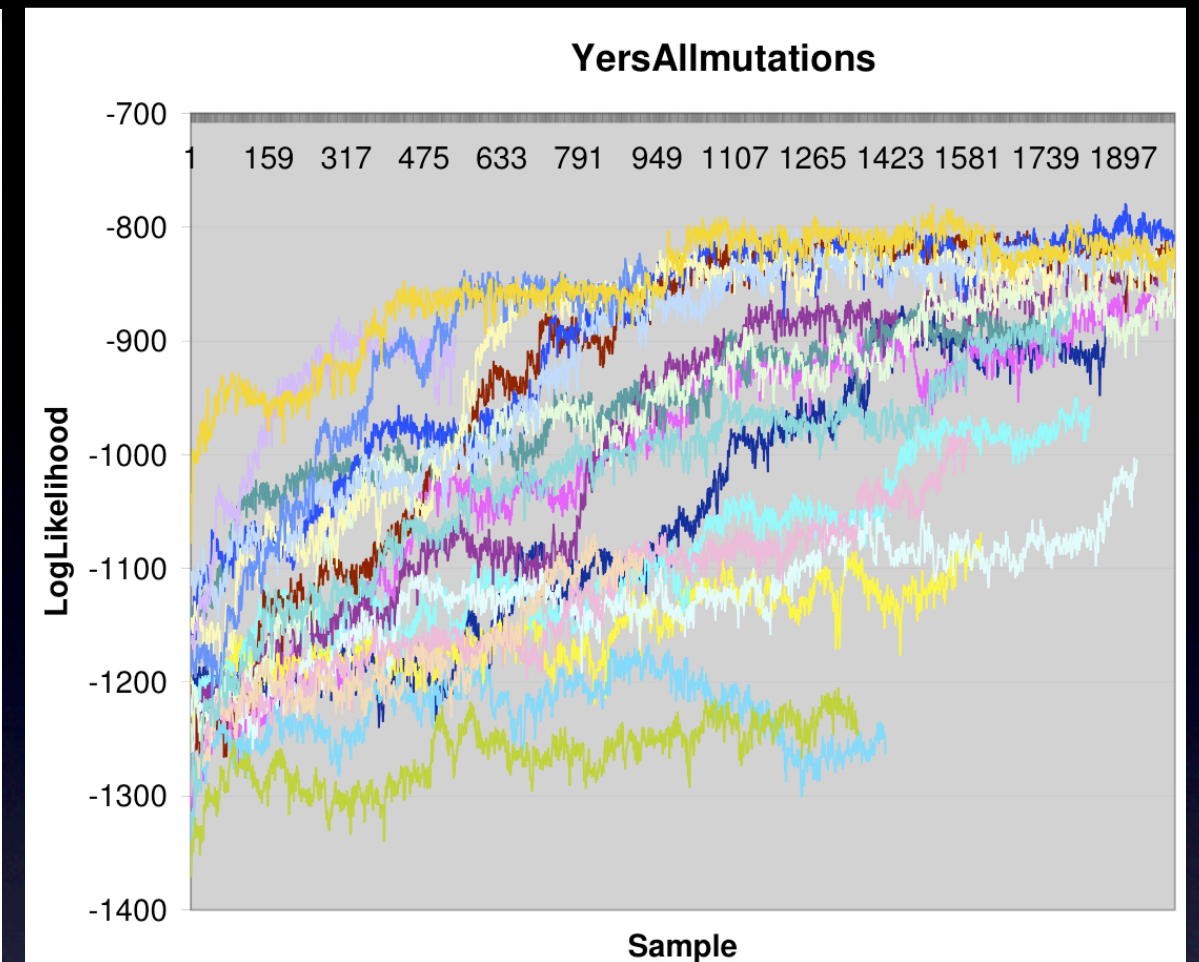
It is worth jumping big

We cannot jump big in discrete, high dimensional distributions.



Myth II.

Small acceptance rate is
caused by bad proposals



Pre-burnin without proposal ratio!

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The 'curse of small backproposal probabilities'.



ParIS: indeed speeded up computation?

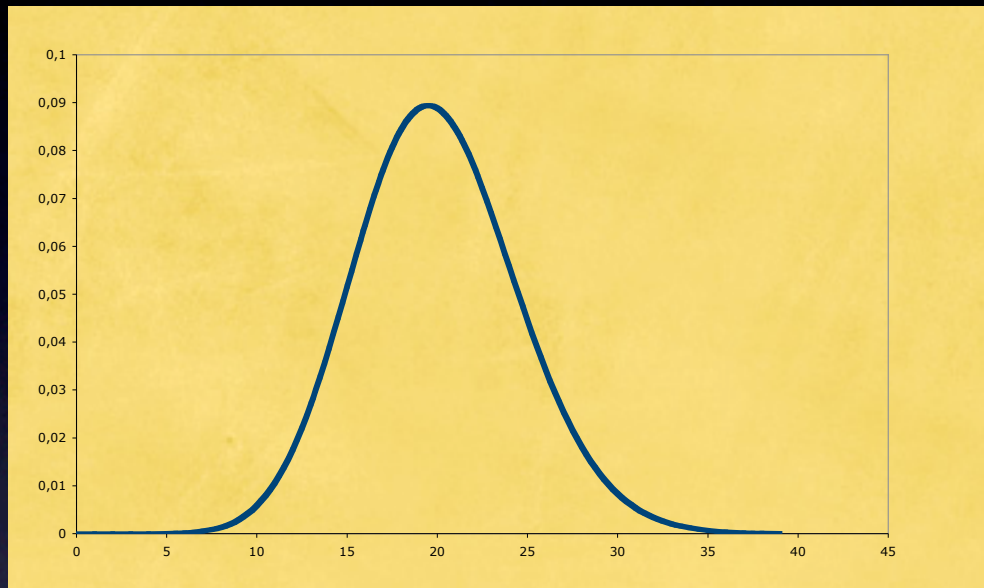
It is enough to calculate proposal probabilities for a 'window' (Green, 1995; Lunter, Miklos, etc, 2005)

$$\pi(X)T(Y|X) =$$

$$\pi(X) \sum_i P(Y, w_i | X) \min \left(1, \frac{\pi(Y)P(X, w_i | Y)}{\pi(X)P(Y, w_i | X)} \right) =$$

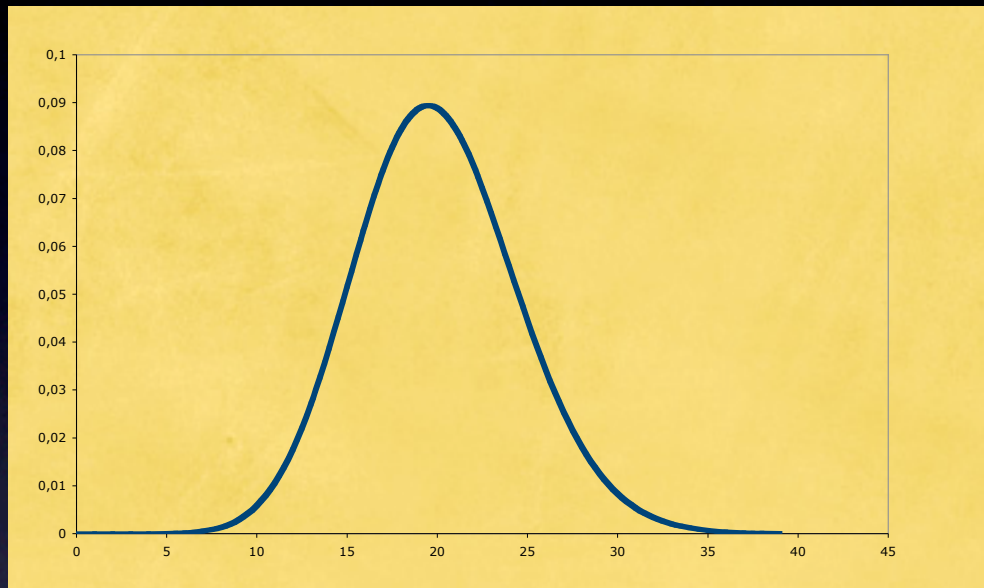
$$\sum_i \min(\pi(X)P(Y, w_i | X), \pi(Y)P(X, w_i | Y))$$

But does it help?



Proposing alignment window lengths
from Poisson distribution

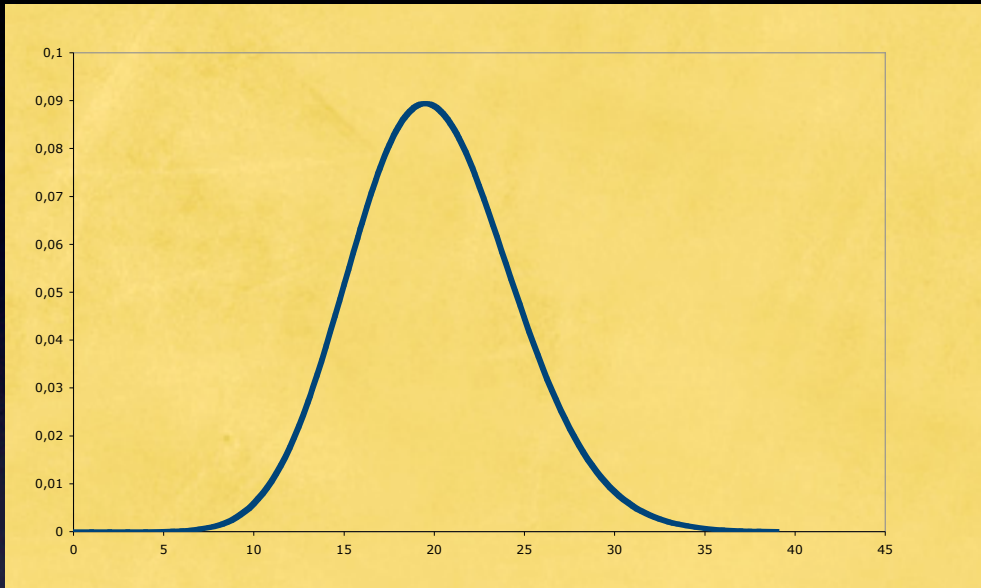
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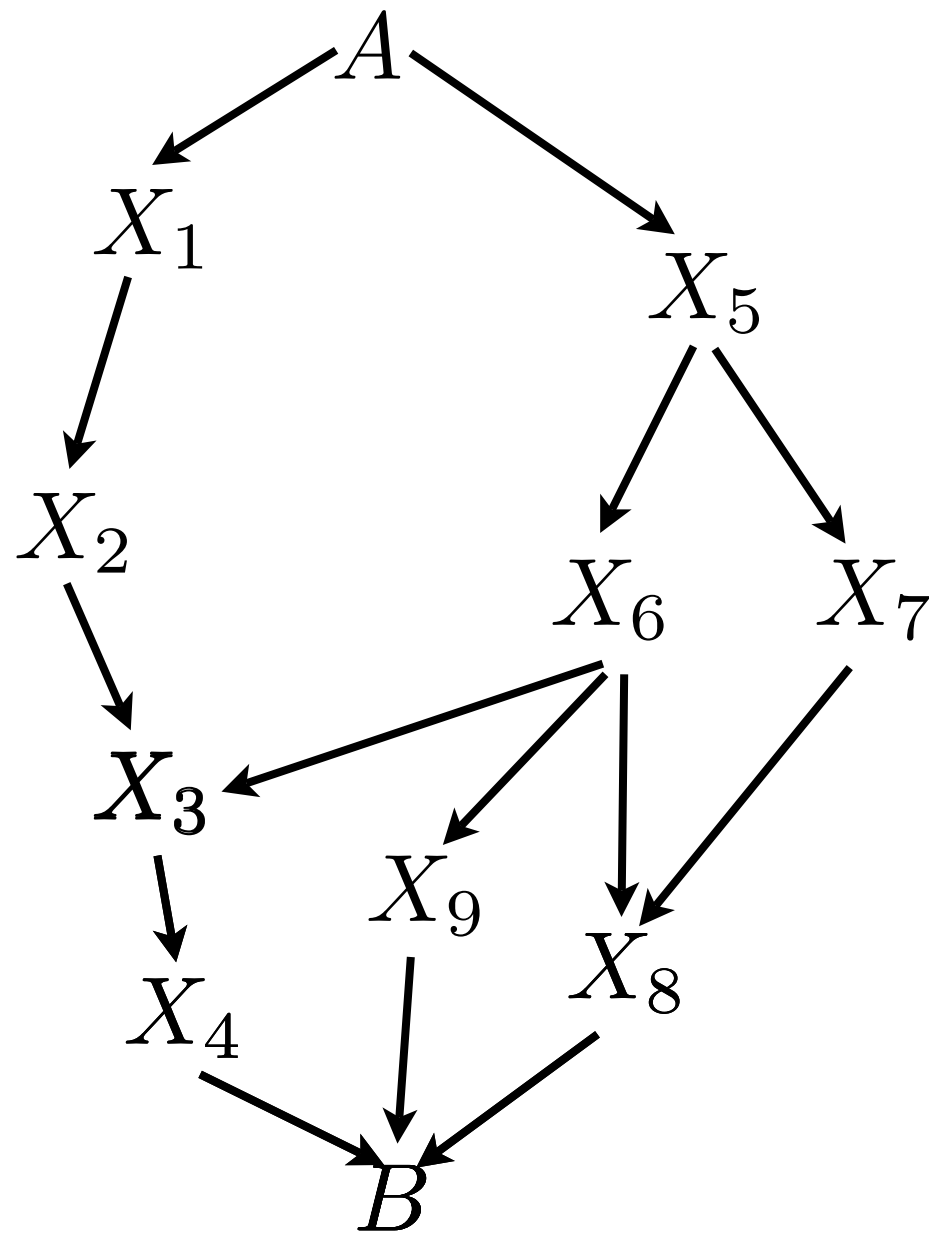


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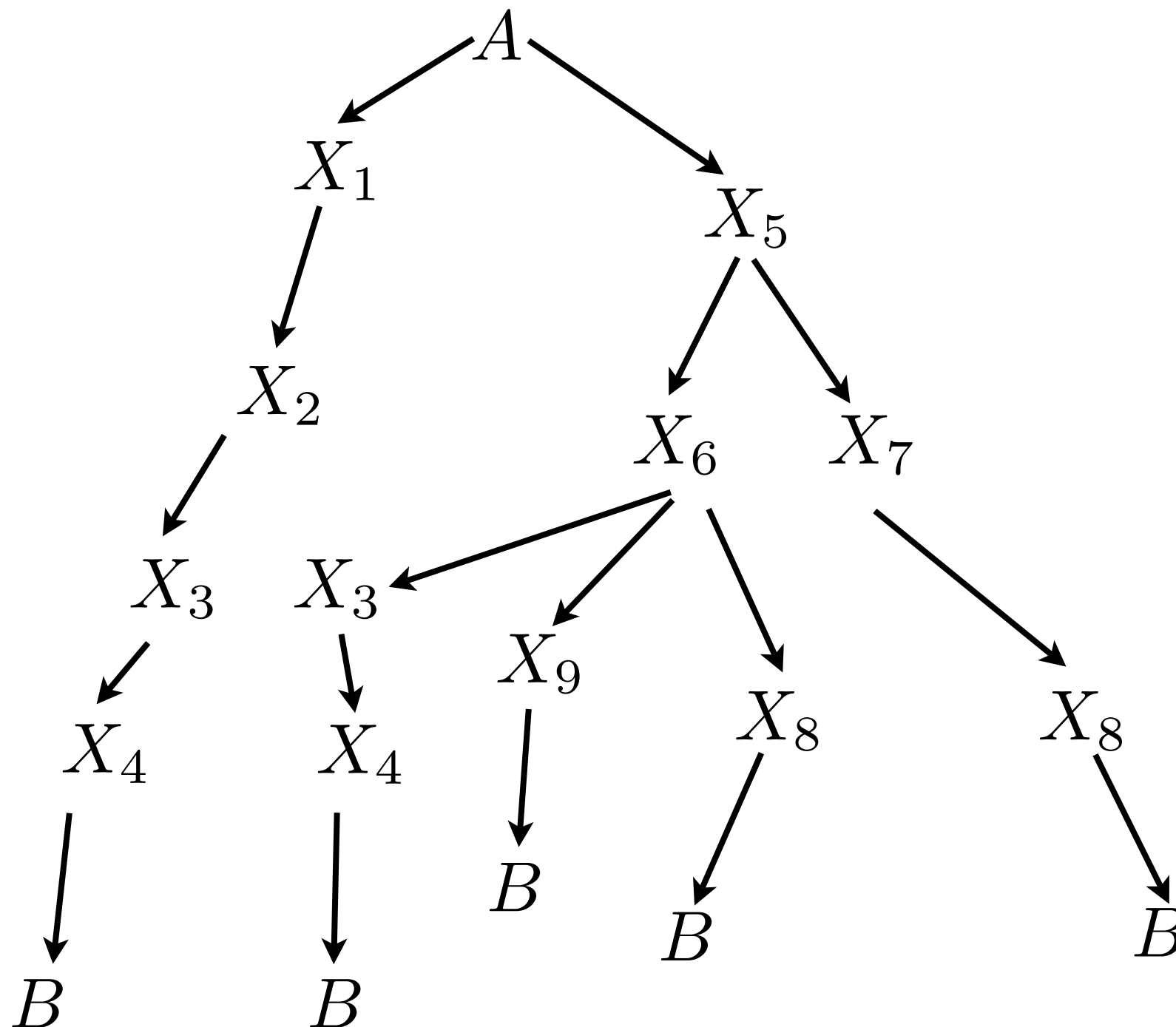
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A mixture distribution helps

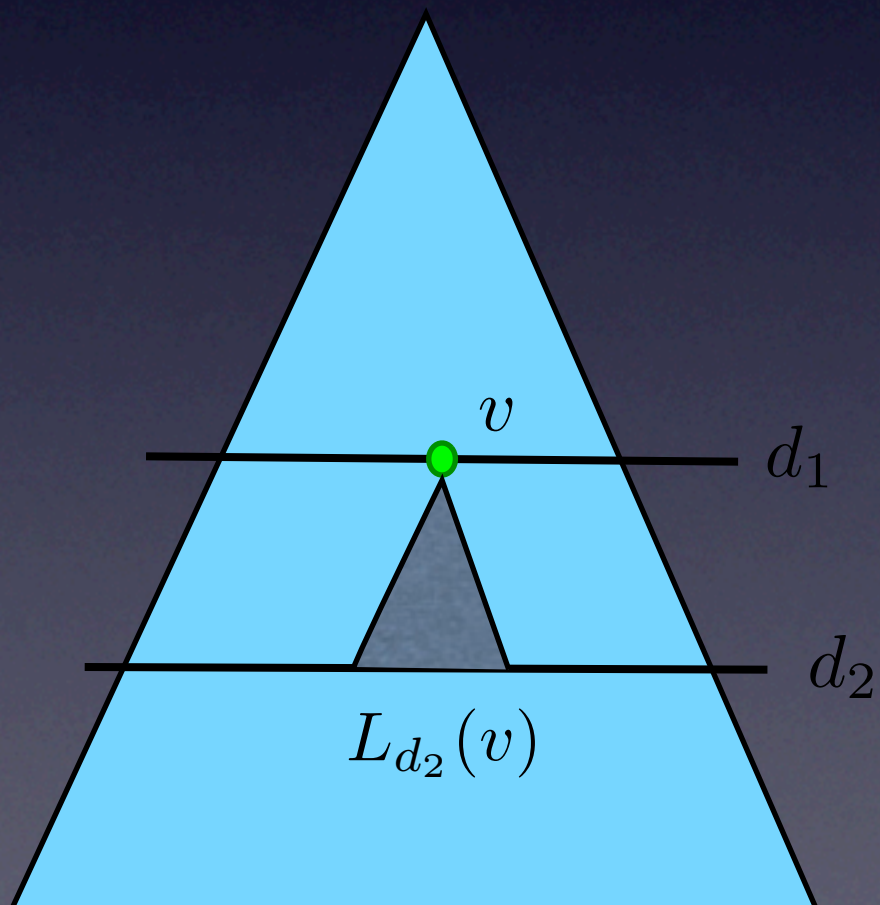
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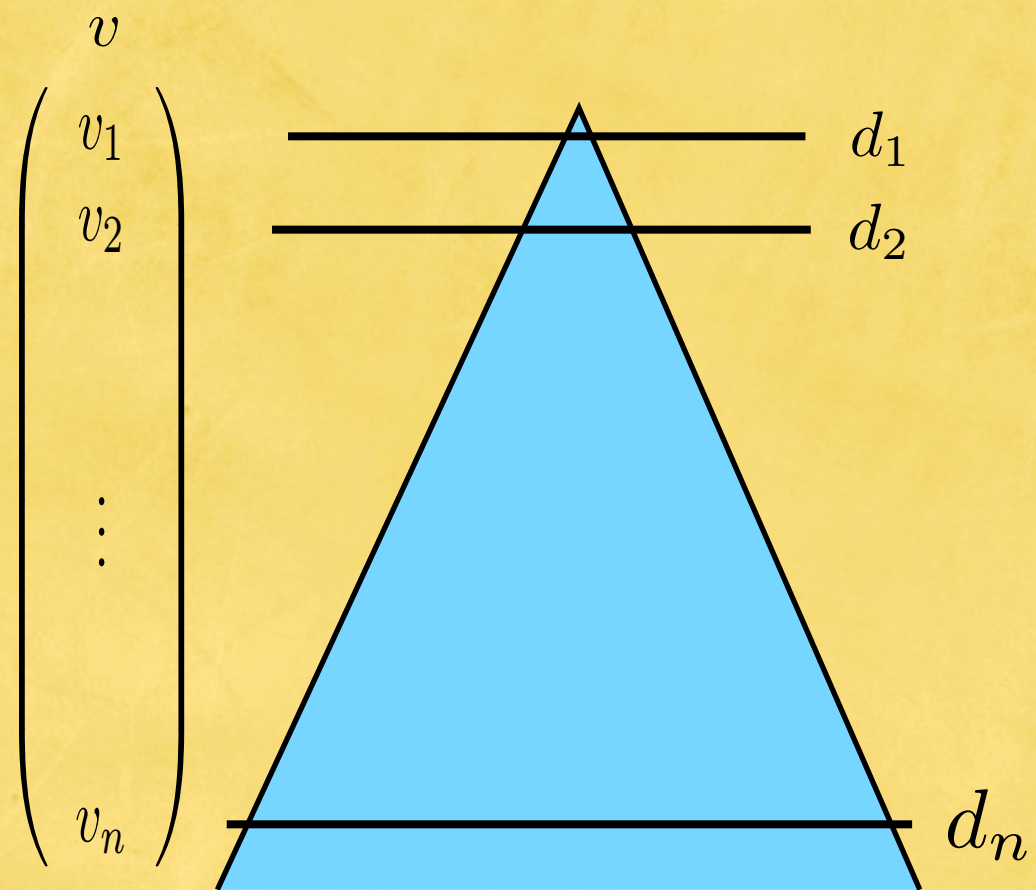
Hiddenness of decision trees



$$|L_{d_2}(v)| \leq k \times \frac{|d_2|}{|d_1|}$$

MCMCMCMC

Model Changing Metropolis Coupled Markov Chain Monte Carlo

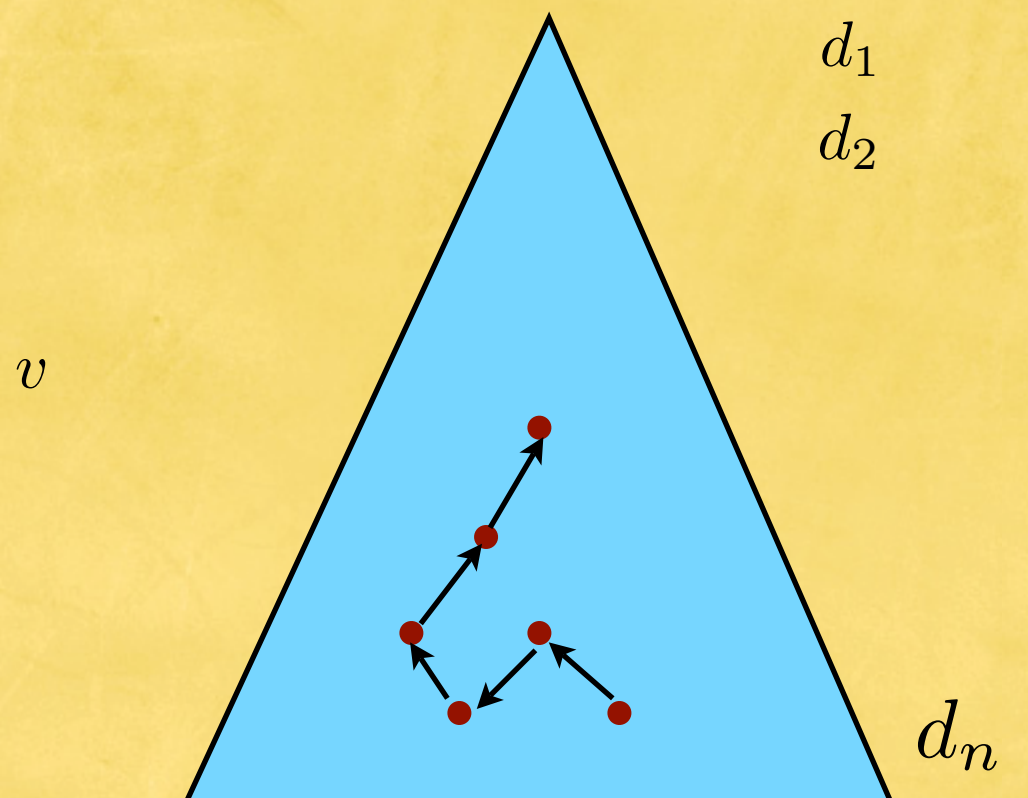


PT-like MSIS

- No need to know the normalizing constant
- Parallelization is possible
- Too many chains?

MCMCMCMC

Model Changing Metropolis Coupled Markov Chain Monte Carlo



ST-like MSIS

- Unknown normalizing constant!
- Parallelization is not possible
- Single chain

Convergence property of MCMCMCMC

For ST-like MSIS

$$\Omega(k) = \tau_{rel} = O(an^2k)$$

For PT-like MSIS

$$\Omega(k) = \tau_{rel} = O(an^{6+\log_2(k)})$$

Where k is the hiddenness rate, 'a' is the smallest transition probability, and n is the depth of the tree

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the length of the signed permutations!

Fortunately, we can add more transition kernels, like
swapping commuting mutations.

Mixing speed is unknown at the moment...

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If the exact answer provides solution for a decision problem and it is a polynomial size of the input

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If the decision problem is NP-complete, then it is unlikely that an FPRAS algorithm exists, unless $BPP \neq P$ or $P=NP$

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$$d_V(\delta_i^T P^n, \pi^T) \leq c(P, \pi) \rho^n$$

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What really matters is the variational distance

$$d_V(\delta_i^T P^n, \pi^T) \leq c(P, \pi) \rho^n$$

$$c(P, \pi) = \min \left\{ \frac{1}{\rho} \left(\frac{p_{ii}}{\pi_i} \right)^{\frac{1}{2}}, \frac{1}{2} \left(\frac{1 - \pi_i}{\pi_i} \right)^{\frac{1}{2}} \right\}$$

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How is this complexity class related to FPRAS?

Conclusions

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I hope I spread non-knowledge significantly today!



Thank you!

